POSITIVE ATTITUDES TOWARD THE OUTGROUP: ADAPTATION AND VALIDATION OF THE ALLOPHILIA SCALE

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The paper presents an Italian validation of the Pittinsky, Rosenthal, and Montoya (2011) Allophilia Scale. The term *allophilia* indicates a positive feeling and attitude of openness toward an outgroup, in this case specifically an ethnic group. Exploratory and confirmatory factor analyses were conducted on a sample of 400 individuals. The sample was randomly split into two parts: an exploratory factor analysis was carried out on the first and a confirmatory factor analysis on the second. Our findings supported the dimensionality of the scale already ascertained with five factors in the American context. The number of items was reduced from 17 to 15 on the basis of our results. A third sample consisting of 180 people was recruited to perform analyses of convergent and discriminant validity, both of which were satisfactory. The analyses gave satisfactory results on the target group examined (the Romanian population residing in Italy).

Key words: Allophilia; Positive inter-group attitudes; Prejudice; Construct validity; Structural equation modeling.

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INTRODUCTION[1]

The proper relationships between positive and negative dimensions of attitudes have recently been the subject of numerous studies that have reached contrasting conclusions. Some research shows that positive and negative aspects of the same attitude are only weakly correlated, suggesting a bidimensionality within the same construct, while other research presents a strong correlation between the two aspects, so much so as to support the hypothesis of a single dimension as best characterizing the attitude.

The relation between pleasure (described as a positive attitude) and displeasure (viewed as a negative attitude) is a longstanding and controversial example of this (Cacioppo & Berntson, 1994; Diener & Iran-Nejad, 1986; Larsen, McGraw, & Cacioppo, 2001; Rozin, Kahneman, Diener, & Schwarz, 1999; Russell & Carroll, 1999). Some researchers, such as Schimmack (2005), consider pleasure and displeasure as mutually exclusive, opposite ends of a single bipolar dimension: "Bipolarity can be defined as a reciprocal relationship between pleasure and displeasure. That is, increases in pleasure imply decreases in displeasure and vice versa. Strong bipolarity would imply a perfectly negative correlation between intensity of pleasure and intensity of displeasure" (Schimmack, 2005, p. 672). In this case, pleasure and displeasure are attitudes that

cannot be experienced simultaneously. However, psychometric evidence does not always support this theory (Russell & Carroll, 1999; Schimmack, Böckenholt, & Reisenzein, 2002).

Other authors, such as Cacioppo and Berntson (1994), consider pleasure and displeasure as distinct attitudes that vary along two separate and mutually exclusive dimensions. Among those who support the latter perspective are proponents of at least two different positions: those who claim that bipolarity is always present (e.g., Schimmack, 2005) and those who maintain, on the other hand, that it appears in ambivalent situations (e.g., Larsen et al., 2001) while, for all non-ambivalent circumstances, experiences of pleasure and displeasure are mutually exclusive (see Table 1).

TABLE 1
Unidimensionality and bidimensionality: Qualifying features

Single bipolar dimension	Bidimensionality
Indicates a single dimension of the construct sub-divided into two poles (positive and negative)	Indicates two dimensions (one positive and one negative).
Increase of the positive pole results in decrease of the negative, and vice versa (If I feel more pleasure, I feel less displeasure)	Both dimensions can increase contemporaneously (If I feel more pleasure, I don't necessarily feel less displeasure).
A very high (negative) correlation exists between the two poles	A low (positive and negative) correlation exists between the two dimensions.
	Two different positions: – Always – Only in ambivalent situations.

Although many studies have attempted to analyze and probe these perspectives empirically, a consensus and shared position has not yet been reached. Recently Schimmack (2001, 2005) proposed a compromise position: pleasure and displeasure are reciprocally correlated, which does not mean that one completely inhibits the other, though.

Measurement issues must also be taken into consideration in order to understand and perhaps help resolve this lively debate. Meddis (1972) demonstrated that response format can influence respondents completing a questionnaire assessing positive/negative polarity. In effect, participants tend to interpret some formats — such as the Likert scale — as reflecting bipolarity and to interpret other response formats — such as intensity scales — as reflecting unipolarity (Gannon & Ostrom, 1996; Russell & Carroll, 1999; Schimmack et al., 2002). This implies that, if positivity and negativity are measured on symmetrical scales, the lowest scores suggest a strong feeling of displeasure, middle scores indicate absence of pleasure/displeasure, and the highest scores indicate presence of strong pleasure. The same response categories carry the same meanings if we use an intensity scale in which the lowest categories indicate absence of pleasure, middle scores indicate moderate pleasure, and the highest scores indicate strong pleasure. In order to overcome this problem, some researchers asked participants to make separate judgments on the presence and intensity of each. In this case, intensity was rated only after participants reported

that an affect was present, through a first rating for presence or absence of each emotion. In this way, research clarified that low responses on an intensity scale refer to low intensity of a present affect and not to strong intensities of an opposing affect (Ekman et al., 1987; Larsen et al., 2001; Reisenzein, 1995). However, this solution has a significant disadvantage: it doubles the number of items to which respondents must answer.

The debate on attitude polarity does not involve only pleasure and displeasure. In the study of ethnic attitudes, the theoretical and conceptual picture has recently been enriched by new theoretical models, with supporting evidence, that present a more complicated view of prejudice. Research on contemporary prejudice has found more subtle, hidden, and refined forms of prejudice which are more multidimensional and multifaceted than previously studied types (Brown, 1997; Dovidio & Gaertner, 1998; Eberhardt & Fiske, 1998; Fiske, 1998; Pettigrew & Meertens, 1995). Katz and Hass (1988) proposed a well-known model of *ambivalent* prejudice characterized by a constant tension between a favorable attitude toward the outgroup and simultaneous unease. According to the theory of ambivalence, people can at the same time experience positive and negative reactions toward ethnic minorities, thus experiencing psychological conflict. The coexistence of negative and positive attitudes leads to an unstable behavioral pattern. An individual can, on the one hand, have attitudes that reflect egalitarian values and emphasize equality and social justice while, on the other, demonstrating a strong defense of traditional values and his or her cultural and national identity. More recently, Fiske and collaborators (Fiske, Cuddy, Glick, & Xu, 2002; Fiske, Gilbert, Fiske, & Lindzey, 1998; Fiske, Xu, Cuddy, & Glick, 1999; Glick, Fiske, & Zanna, 2001) reinforced this general theoretical picture in their research by suggesting that, in the majority of cases, social relations are multifaceted and polyhedral, that is, underpinned by feelings of differing tone based on ambivalence. The search for a proper understanding of attitudes toward ethnic others — particularly when they are rightly seen as uni-dimensional or multi-dimensional — continues.

With respect to ethnic attitudes, Pittinsky, Rosenthal, and Montoya (2011) recently argued for the investigation of positive ethnic attitudes independently, that is without inferring them from their opposites. In fact, they suggested that positive intergroup attitudes can be identified and measured and predict positive behavior toward outgroups better than negative attitudes (Pittinsky et al., 2011; Pittinsky, Rosenthal, & Montoya, in press). Other evidence supports their claim and findings. A small but growing number of studies suggests that positive intergroup attitudes predict positive intergroup orientations generally better than negative intergroup attitudes. Such research reinforces both the view that positive attitudes toward outgroups should be studied autonomously and the view that they must be measured independently of negative attitudes. Stangor, Sullivan, and Ford (1991) found that, when overall ratings were framed as "favorability" ratings (vs. "unfavorability" ratings), positive affect toward outgroups accounted for significantly more variance than negative affect. Similarly, in a study of attitudes toward Asian Americans, Ho and Jackson (2001) observed that positive intergroup attitudes toward Asian Americans were specifically related to positive dependent variables (i.e., admiration and acceptance of positive stereotypes), whereas negative attitudes were specifically related to negative dependent variables (i.e., hostility, acceptance of negative stereotypes, and greater social distance). Other research suggested that positive emotions are more closely linked to positive intergroup behaviors. For instance, Dijker (1987; Dijker, Koomen, van den Heuvel, & Frijda, 1996) found that positive emo-

tions (e.g., enthusiasm, happyness) toward ethnic minority groups were correlated with contact with members of those groups. Also, Tam et al. (2007) found that such positive emotions were independent of negative emotions in the prediction of behavior. Among political factions in Northern Ireland, positive intergroup emotions only predicted approach behaviors, whereas negative intergroup emotions predicted aggressive and avoidant behaviors (Tam et al., 2007). The picture, however, is not entirely clear. Research on intergroup contact, for example, provides evidence that negative emotions, such as anxiety, can also be pivotal in predicting positive behavior or, at least, positive behavioral intentions, such as desire for future contact (e.g., Levin, Van Laar, & Sidanius, 2003; Plant & Devine, 2003).

Taken together, while the ultimate degree of independence of positive and negative intergroup attitudes awaits further study, the research on positive intergroup attitudes to date supports the hypothesis that it is crucial to consider such attitudes independently of negative attitudes, particularly when attempting to understand positive intergroup behaviors (Pittinsky et al., 2011, in press).

AIMS OF THE STUDY

Based on the research described above, an understanding of positive attitudes appears to require a direct measurement of them, rather than deducing their presence from low levels of negative attitudes. Pittinsky et al. (2011), consistently with this position, investigated such positive outgroup attitudes without conceiving them as simple lack of prejudice. They proposed the Allophilia construct (Pittinsky, 2010) to refer to positive attitudes toward a given ethnic outgroup. The term *allophilia*, of Greek etymology, literally means "liking for others" and indicates an individual feeling toward an outgroup not reducible to mere tolerance. According to Pittinsky et al. (2011), and as demonstrated in their studies, Allophilia (a) is distinct from qualified forms of liking, (b) is not simply the opposite of negative prejudices and racism, and (c) plays a distinct role in intergroup relations, adding variance to the prediction of positive intergroup behaviors independent of negative intergroup attitudes, positive but qualified intergroup attitudes, general prosocial orientations, group identification, categorization, and norms. Allophilia has five cardinal factors:

– *Affection*, defined as a positive feeling experienced toward an outgroup (sample scale item, "In general, I have positive attitudes about African Americans");
– *Comfort*, that is, feeling at ease with outgroup members ("I am at ease around African Americans");
– *Kinship*, in the sense of "sharing something," experiencing a sense of belonging with the outgroup ("I feel a sense of kinship with African Americans");
– *Engagement,* conceived as seeking contact with the outgroup ("I am truly interested in understanding the points of view of African Americans");
– *Enthusiasm*, that is, having a favorable impression of outgroup members ("I am favorably impressed by African Americans").

These factors are compatible with the tripartite model of attitudes (e.g., Katz & Stotland, 1959; Rosenberg & Hovland, 1960). Affection, Comfort, Kinship, and Enthusiasm consist of af-

fective evaluations. Engagement consists of behavioural evaluations. As noted in the original paper (Pittinsky et al., 2011), cognitive evaluations were excluded from the scale early in the process.

The aim of the present work was to make an initial contribution toward the validation of the allophilia construct, which had been validated in the United States, in the Italian setting. The key feature of this construct is that it theoretically and empirically treats a positive intergroup attitude as a distinct experience, and not simply as the opposite of a negative attitude toward an outgroup.

The present research comprises three studies. The first is an exploratory factor analysis conducted on half of the group of participants, aimed to verify several psychometric properties of the Allophilia Scale (Pittinsky et al., 2011). Following the information obtained in the first study, a second study carried out a confirmatory factor analysis on the remaining half of the sample, aimed to corroborate the factorial structure of the scale. Finally, a third study, performed on a third group, sought to analyze the convergent and divergent validity between the allophilia construct and other constructs considered.

The choice was made to maintain, as much as possible, procedures and methodologies already adopted by Pittinsky et al. (2011) in the American context.

STUDIES 1 AND 2

Method

Participants and Procedures

A total of 400 individuals took part in the first two studies: 126 non-student adults (31.5%) and 274 university students (68.5%), 89.7% of whom were enrolled in the Department of Economics at the Catholic University in Milan while 10.3% were enrolled in the Departments of Psychology and Developmental Sciences. Total mean age was 24.31 years (SD = 8.04, Range 20-68), with 43.7% of participants being male and 56.3% female. No statistically significant differences emerged for age with respect to participant gender $t(394) = -1.901, p = .058$.

To guarantee the random distribution of respondents in the two studies, the sample was split into two equal parts by means of the SPSS command, *Random Sample of Cases* (Byrne 1994; Huizingh, 2007; Reis & Judd, 2000), with the request to extract 50% of the sample. This resulted in two datasets. The first was used for the exploratory part of the analysis (Study 1) and the second for the confirmatory analysis (Study 2).

For Study 1, SPSS extracted a randomized subsample of 209 participants from the database 65.1% of whom were students and 34.9% employed. The mean age was 24.79 years (SD = 9.02, Range 20-68) with 56.1% of respondents being female and 43.9% male. For Study 2, the confirmatory study, the second half of the sample was used. The subsample extracted by SPSS consisted of 191 participants. Of these, 72.3% were students and 27.7 were employed, with 55.5% being female and 44.0% male. The average age was 23.77 years (SD = 6.81, Range 20-58).

Most participants lived in the Italian region of Lombardy (84.9%); 46.3% lived in province capitals and the remaining 53.7% in smaller urban centers.

Recruitment of student participants took place during lessons. Recruitment of adult participants was through snowball sampling by requesting students to also have adults fill out the questionnaire given to them. Everyone was informed from the beginning that participation was voluntary, without any form of compensation, and that all the data would be treated confidentially and used only for research purposes. After completing the study, participants were briefly informed about the aim of the study. The decision to include both student and adult participants was made with the goal of investigating the reliability of the instrument for both generations. This would help find out, for example, whether the scale can be used in research on families, which requires a comparison between parents' and offspring's mean scores.

Materials

An Italian version of the Allophilia Scale was created by two independent translators. An Italian psychologist proficient in English carried out the translation from English into Italian; an American native speaker with a good knowledge of psychology translated the scale back into English. The Kappa coefficient for the agreement between the translators was .89.

The Allophilia Scale was composed of 17 items. Responses were given on a 6-step scale ranging from 1 (*strongly disagree*) to 6 (*strongly agree*). Low scores thus indicated a respondent's low levels of allophilia for a group while high scores indicated the opposite. Each item loaded on only one of the five factors.

The target population of the original American version was African-Americans, though the authors showed that it worked well for many different groups in the United States. Since there is no analogous population in Italy in terms of historical and socio-cultural characteristics, the Romanian population was chosen instead. Romanians have long been the subject of considerable media attention and are often recognized as being a target of prejudice and discrimination. This was especially true at the time of data collection due to several concurrent events in the news. Participants were asked to complete a questionnaire which included the Allophilia Scale preceded by several demographic variables, such as age, gender, city of residence, profession (if employed), and area of study (if a student).

STUDY 1

Data Analyses

Following Gerbing and Hamilton (1996), who reasoned that exploratory factor analysis can be used prior to any analysis technique aimed to confirm hypotheses on data structure, an exploratory factor analysis was carried out using SPSS software V. 15.0. The extraction method employed was Principal Axis Factoring and Promax Rotation. Although the Oblimin rotation is the most widely used in the literature, and there is a substantial convergence of results (Fabrigar, Wegener, MacCallum, & Strahan, 1999), Promax appeared to offer a solution that would be more replicable than an Oblimin rotation (Rennie, 1997).

Results

In our sample, both skewness and kurtosis for the majority of items fell between –1.00 and +1.00.

The Kaiser-Meyer-Olkin Measure of Sampling Adequacy was found to be above .7 (.95), and Bartlett's Test of Sphericity was significant. All communalities of items had satisfactory values (between .50 and .86; see Table 2). Five factors were submitted to SPSS, as called for in the original scale, for a total explained variance of 78.15%.

TABLE 2
Communalities

	Initial	Extraction
1. In general, I have positive attitudes about Romanians	.73	.80
2. I respect Romanians	.60	.65
3. I like Romanians	.78	.76
4. I feel positively toward Romanians	.83	.83
5. I am at ease around Romanians	.80	.84
6. I am comfortable when I hang out with Romanians	.78	.83
7. I feel like I can be myself around Romanians	.49	.50
8. I feel a sense of belonging with Romanians	.81	.86
9. I feel a kinship with Romanians	.79	.86
10. I would like to be more like Romanians	.67	.70
11. I am truly interested in understanding the point of view of Romanians	.75	.80
12. I am motivated to get to know Romanians better	.80	.85
13. To enrich my life, I would try and make more friends who are Romanian	.73	.72
14. I am interested in hearing about the experiences of Romanians	.73	.81
15. I am impressed by Romanians	.75	.77
16. I feel inspired by Romanians	.75	.80
17. I am enthusiastic about Romanians	.77	.82

Note. Extraction method: Principal axis factoring.

The Pattern Matrix showed a net item distribution on five factors, for all items (Table 3).

TABLE 3
The Pattern Matrix

	Affection	Comfort	Kinship	Engagement	Enthusiasm
1. In general, I have positive attitudes about Romanians	**.916**	−.020	.070	−.023	−.042
2. I respect Romanians	**.876**	−.037	−.157	.175	−.098
3. I like Romanians	**.492**	.156	.051	−.043	.295
4. I feel positively toward Romanians	**.636**	.105	.168	−.108	.172
5. I am at ease around Romanians	.327	**.533**	.091	−.034	.084
6. I am comfortable when I hang out with Romanians	.033	**.623**	.114	.020	.204
7. I feel like I can be myself around Romanians	.260	**.513**	−.036	.131	−.101
8. I feel a sense of belonging with Romanians	.020	.103	**.907**	−.015	−.156
9. I feel a kinship with Romanians	.009	.078	**.916**	.017	−.070
10. I would like to be more like Romanians	−.065	−.165	**.805**	.008	.223
11. I am truly interested in understanding the point of view of Romanians	.002	.148	−.015	**.901**	−.132
12. I am motivated to get to know Romanians better	−.041	.186	−.027	**.787**	.084
13. To enrich my life, I would try and make more friends who are Romanian	.189	−.075	.284	**.433**	.141
14. I am interested in hearing about the experiences of Romanians	−.067	−.169	.031	**.877**	.083
15. I am impressed by Romanians	.133	−.029	.132	.096	**.619**
16. I feel inspired by Romanians	.115	.068	.013	−.037	**.769**
17. I am enthusiastic about Romanians	−.101	.027	.231	.102	**.697**

Cronbach's Alpha was high for all factors ($\alpha = .91$ for Affection; $\alpha = .86$ for Comfort; $\alpha = .91$ for Kinship; $\alpha = .92$ for Engagement; $\alpha = .91$ for Enthusiasm).

STUDY 2

Data Analyses

The dimensionality of the Allophilia Scale was evaluated by testing measurement and structural equation models. Goodness-of-fit was checked using several indexes simultaneously (Bollen, 1989). Two indexes were χ^2 and the ratio between χ^2 and degree of freedom (χ^2/df). A model fits the data well when χ^2 is non-significant. An acceptable ratio for χ^2/df should be less than 3.0 (Hair, Black, Babin, Anderson, & Tatham, 2006). However, the chi-square statistic is sensitive to the sample size. For this reason, we adopted further fit indexes that were less sensitive to the sample size: CFI, SRMR, and RMSEA. CFI values equal to or above .95 were considered acceptable (Hu & Bentler, 1997, 1999). As to SRMR, Hu and Bentler (1999) suggested that the fit is satisfactory for values of .08 or below. RMSEA values equal to or smaller than .08 were considered satisfactory.

Results

To examine the dimensionality of the Allophilia Scale, we applied CFA (LISREL 8.54; Jöreskog & Sörbom, 2004). A model with five dimensions and 17 items was evaluated. It showed a good fit: $\chi^2(109) = 267.53$, $p \cong .000$, $\chi^2/df = 2.45$; CFI = .98, SRMR = .05, RMSEA = .08 (see Figure 1). All the factor loading values were significant or very close to significance (.06 and .07) and with values higher than .67.
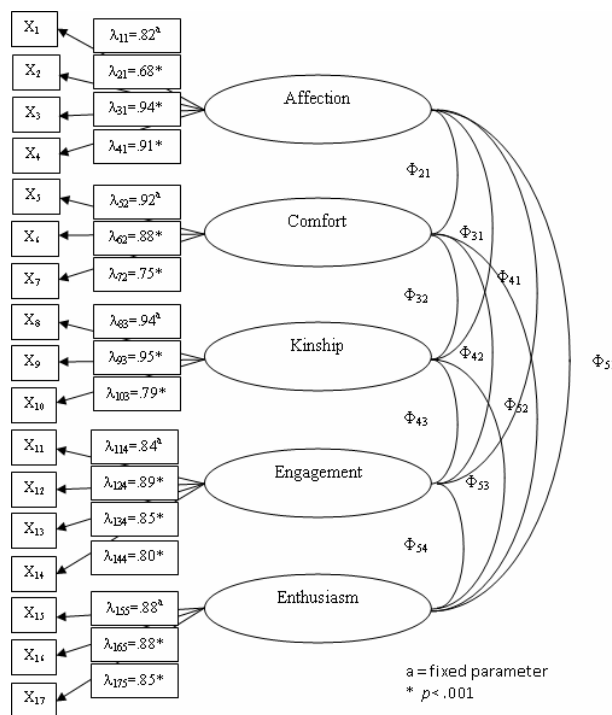

FIGURE 1
The five-factor model for Allophilia.

Construct Reliability and Average Variance Extracted (AVE) also gave satisfactory results (respectively, .86 and .68 for Affection, .89 and .74 for Comfort, .91 and .77 for Kinship, .89 and .73 for Engagement, .89 and .74 for Enthusiasm).

As hypothesized, we found a positive and significant correlation between all five subdimensions (see Table 4).

TABLE 4
Correlation (Φ coefficients) between subdimensions

	Affection	Comfort	Kinship	Engagement	Enthusiasm
Affection	–				
Comfort	.92**	–			
Kinship	.65**	.75**	–		
Engagement	.67**	.62**	.61**	–	
Enthusiasm	.82**	.85**	.79**	.77**	–

** $p < .001$.

Since Affection and Comfort turned out to be strongly correlated ($r = .92$), an alternative four-dimensional model was tested in which they were considered as a single factor. This model showed a medium fit: $\chi^2(113) = 319.89$, $p < .000$, $\chi^2/df = 2.683$; CFI = .97, SRMR = .06, RMSEA = .10. In both statistical and conceptual terms, the former model is preferable, though.

As further confirmation of the dimensionality of the factorial structure, we wanted to test an alternative model with 17 items and a single latent factor. This model provided a poor fit: $\chi^2(119) = 869.57$, $p < .000$, $\chi^2/df = 7.29$; CFI = .90, SRMR = .09, RMSEA = .09.

CONCLUSION
STUDIES 1 AND 2

The aim of the factor analysis presented here was to investigate the dimensionality of an Italian version of the Allophilia Scale, first reported by Pittinsky et al. (2011). Preliminary analyses revealed that the factorial structure of the scale remains unchanged from that in the American context, despite the change in language and historical and socio-cultural characteristics than those of African Americans. The articulation of the five dimensions investigated remains the same, as well as the factor loading of items.

STUDY 3

Analyses of Convergent and Discriminant Validity

Participants and Procedures

One hundred eighty people participated in Study 3: 38 males (21.1%) and 142 females (78.9%), with a mean age of 26.01 years ($SD = 8.81$, Range 19-60). The majority of participants

(89.8%) lived in Lombardy. The hometown of 41.7% of respondents were province capitals while 58.3% lived in smaller urban centers. Exactly the same percentages represent students in the Department of Developmental Sciences at the Catholic University (41.7%) and employed participants (58.3%).

Materials

In order to verify the convergent validity of the Allophilia Scale when no other scales exist to measure positive attitudes toward an outgroup, a comparison with a scale measuring negative prejudice was carried out, following the example of the authors of the American version (Pittinsky et al., 2011). The underlying hypothesis was that the Allophilia Scale should have a reasonably strong negative correlation with a prejudice construct, yet not as high as to suggest that it is its mere opposite, as previously discussed.

In order to validate the original version of the Allophilia Scale, the authors used a variety of scales that could not be used in the present study because they have not been adapted to and validated in the Italian context. To date, there are still no validated instruments in Italian that measure positive attitudes toward an ethnic outgroup. Such instruments would provid valuable information as to the convergent validity. The most frequently used scale validated in Italian aimed to measure negative ethnic prejudice is the Subtle and Blatant Prejudice Scale by Pettigrew and Meertens (1995) in the Italian version by Arcuri and Boca (renamed Latent and Manifest Prejudice Scale, 1996). A comparison was also carried out with other measurement scales traditionally correlated with prejudice: the Social Dominance Orientation Scale in the version by Aiello, Chirumbolo, Leone, and Pratto (2005) and the Ethnocentric Scale by Aiello and Areni (1998).

Finally, in order to investigate discriminant validity, a comparison was made with Rosenberg's (1965) Self-Esteem Scale, using Prezza, Trombaccia, and Armento (1997) Italian version, and with Caprara, Capanna, Steca, and Paciello's (2005) Prosociality Scale, thus following the authors' example in validating the scale's English-language version. All five scales described have response formats based on agreement-disagreement poles. Table 5 shows a sample item for each scale and the respective Cronbach's α.

Convergent Validity

The Allophilia Scale demonstrated a high negative and statistically significant correlation with the Manifest Prejudice dimension ($r = -.82$, $p < .001$), while the correlation with Latent Prejudice, still negative, was decidedly smaller, albeit statistically significant ($r = -.38$, $p < .001$). When subdividing the Allophilia Scale into its five dimensions, as called for by the authors, a substantial difference emerged with respect to correlations with the Prejudice Scale. Indeed, while Manifest Prejudice had high negative correlations (from a minimum of $-.55$ for Kinship to a maximum of $-.78$ for Affection), Latent Prejudice had decidedly lower, though statistically significant, correlations (from a minimum of $-.28$ to a maximum of $-.39$). As regards the Social

TABLE 5
Sample items and Cronbach's alphas for the scales used

	Sample item	Cronbach's α
Manifest Prejudice	Italians and Romanians will never be able to feel at ease with each other, even when they become friends	.78
Latent Prejudice	It would be better for Romanians who live in our country to avoid places in which their presence is not welcome	.72
Social Dominance Orientation Scale	It is undeniable that some groups of people in our society are more respected	.89
Ethnocentrism	People often worry about the lack of comprehension that Italians show for those whose culture is different, but it seems that no one is asking whether the latter are showing comprehension towards Italians	.90
Self-Esteem Scale	I think that I am worth at least as much as other people	.60
Prosociality Scale	I willingly help my classmates and friends in carrying out their activities	.89

Dominance Orientation Scale, statistically significant and negative correlations emerged, for both the Allophilia construct ($r = -.51$ $p < .001$) and all five dimensions composing it. Also the correlations between Ethnocentrism and Allophilia — considered in its globality ($r = -.69$, $p < .001$) as well as in its five dimensions — were high, statistically significant, and negative.

The correlations of Social Dominance Orientation Scale and Ethnocentrism Scale did not, however, obtain high scores, contrary to that of Manifest Prejudice, although all were of medium-to-large magnitude. The correlations with these scales and the single subdimensions of Allophilia are reported in Table 6.

TABLE 6
Pearson correlation between Manifest and Latent Prejudice, Social Dominance Orientation, and Ethnocentrism and Allophilia

	Allophilia	Affection	Comfort	Kinship	Engagement	Enthusiasm
Manifest Prejudice	–.82**	–.80**	–.77**	–.55**	–.72**	–.70**
Latent Prejudice	–.38**	–.37**	–.28**	–.27**	–.30**	–.39**
Social Dominance	–.51**	–.53**	–.47**	–.33**	–.49**	–.38**
Ethnocentrism	–.69**	–.65**	–.60**	–.53**	–.60**	–.60**

** $p < .001$.

Discriminant Validity

The Allophilia Scale turned out to have a very low, albeit statistically significant, correlation with the Prosociality Scale. All five dimensions of the Allophilia Scale had very low correla-

tions as well, although Kinship and Engagement were statistically significant ($r = .19$, $p < .05$; $r = .22$, $p < .01$, respectively). Likewise, as regards the Self-Esteem Scale, correlations were null and not statistically significant (Table 7).

TABLE 7
Correlation between Prosociality, Self Esteem, and Allophilia

	Allophilia	Affection	Comfort	Kinship	Engagement	Enthusiasm
Prosociality	.15*	.06	.11	.19*	.22**	.13
Self Esteem	.04	.05	.05	.06	.01	.02

*$p < .05$; ** $p < .01$.

CONCLUSIVE REMARKS
STUDY 3

The aim of Study 3 was to investigate the convergent and discriminant validity of the Allophilia Scale when compared to other constructs which should be positively related (in the first case) and negatively related (in the second).

With regard to convergent validity, to date no other scales have been developed to measure the Allophilia construct or overall positive attitudes toward an ethnic outgroup. Thus, scales aimed to measure analogous constructs and validated in the Italian construct were chosen: the Latent and Manifest Prejudice Scale, the Social Dominance Orientation Scale, and the Ethnocentric Scale. All of these investigated dimensions turned out, as expected, to have negative, statistically significant correlations of medium-to-large magnitude. The highest correlations were between the Allophilia Scale and Manifest Prejudice, which thus seemed to be the most conceptually overlapping constructs. The weaker, although statistically significant, correlation between Allophilia and Latent Prejudice was particularly interesting. This result seems to suggest that the Manifest Prejudice Scale and the Allophilia Scale are substantially opposite to each other, while Latent Prejudice and Allophilia are two distinct, albeit correlated, aspects. Another interpretation, however, as proposed by Pittinsky et al., 2011, could be that Allophilia and Prejudice differ in their predictive validity with respect to positive or negative behaviors toward the outgroup.

The Social Dominance and Ethnocentrism Scales, traditionally used together with the prejudice scales, correlated in a negative and statistically significant way with the Allophilia Scale, although with a weaker link when compared to that found for Manifest Prejudice. These results provided support for the initial hypothesis that, since the Allophilia Scale has a middle negative correlation with the investigated constructs, Allophilia should not be considered their mere opposite, a perspective that was discussed and argued for earlier. It is not possible to carry out a comparison with the original validation as the authors used only scales aimed to detect positive attitudes toward the outgroup in order to test the convergent validity (e.g., The Pro-Latino Scale by Katz & Hass, 1988; Universal Orientation Scale by Philips & Ziller, 1997).

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In conclusion, the Allophilia Scale turned out to be valid as far as convergence is concerned, in particular compared to the Latent Prejudice Scale, the Social Dominance Scale, and the Ethnocentric Scale. It is not possible to infer further with regards to the Manifest Prejudice Scale.

To assess the scale discriminant validity, the Self-Esteem Scale and the Prosociality Scale were examined. No statistically significant correlations emerged between Allophilia (globally and in the five dimensions) and Self Esteem, as was also the case in the original American version ($r = .08$, *ns*). Allophilia and Self Esteem thus did not measure analogous constructs. Low but statistically significant correlations were found between Allophilia and the specific factors of Kinship and Engagement. In the original version, the authors also found a statistically significant correlation between global Allophilia and Kinship and Engagement, of greater magnitude than that found in the present study ($r = .26$, $p < .001$). They did not provide any explanations for this connection, though.

In summary, the Allophilia Scale turned out to have good discriminant validity. It clearly differentiated from the Self-Esteem Scale and correlated only for some dimensions, and with a very low magnitude, with the Prosociality Scale.

DISCUSSION AND CONCLUSION

The present work sought to contribute to the validation of the Allophilia Scale in an Italian context. Our findings essentially confirmed the validity of the Italian version. We initially tested the scale by using an exploratory factor analysis with 17 items and five subdimensions. All of the items showed satisfactory factor loadings in relation to the respective factors. communalities are high and the total percentage of variance explained is very strong. Subsequently, a confirmative factor analysis was carried out, keeping the five 17-item dimensions. The fit indexes were satisfactory, suggesting a good applicability of the scale to the Italian context. The final version thus confirms the factorial structure of the original. Nevertheless, the dimensions of Affection and Comfort are strongly correlated (.92), suggesting a possible dimensionality reduced to four factors. However, this model did not provide fit indexes equivalent to the five-dimension model. A model with a latent factor and 17 items was also tested, but did not provide satisfactory fit indexes.

Convergent and discriminant validities were tested on the 17-item scale. The results deserve some consideration as decidedly high and statistically significant negative correlations emerge in the Manifest Prejudice Scale. This finding raises many questions. Firstly, about the high correlation between the Allophilia Scale and the Manifest Prejudice Scale. However, no similar finding was observed for Latent Prejudice, which, instead, appears to confirm the hypothesis that it measures a construct distinct from that of Allophilia. At present, it is not possible to establish whether Allophilia and Manifest Prejudice are opposites or two different constructs. Future examination of their respective predictive validity would likely be important to clarify this issue. In fact, according to research by the scale's authors, Allophilia plays a distinct role in intergroup relations, adding variance to the prediction of positive intergroup behaviors independent of negative intergroup attitudes (Pittinsky et al., in press). As previously pointed out, however, our interest did not lie in unraveling such a complex knot, which calls for subsequent research and analysis for clarification.

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The Allophilia Scale produced satisfactory results in its differentiation form the Self-Esteem Scale and Prosociality Scale: it correlates only to some dimensions of these and at low levels. On the whole, results are satisfactory. They suggest that the Allophilia construct articulated in five factors is well suited to the Italian context and, in particular, to the attitudes toward Romanians as a target population. Moreover, the simplicity and brevity of the scale's items make this instrument easy to administer and to understand. Some limitations to the current research should be pointed out, though, most notably the geographical location of the sample. Thus, in future research, a more representative sample of the Italian population should be considered, including greater diversity in participant socio-economic, cultural, and ethnic background. Secondly, predictive validity — which, according to Pittinsky et al. (2011), explains the higher percentage of variance in the prejudice scales — has not yet been investigated; for example, in predicting charitable behaviors toward associations supporting immigrants (Pittinsky et al., in press).

In conclusion, the present study can be considered as a solid first contribution toward the adaptation of the Allophilia Scale to the Romanian population in Italy. Analyses parallel to those reported here provide analogous findings for other target populations — for example, generic "immigrants" (Alfieri, Marta, & Lanz, 2011) — which appears to testify to the scale's flexibility and thus to its use in multiple contexts. In addition, we should not underestimate the possibility that, as the etymological root of the term "allophilia" suggests, the target group does not necessarily have to be an ethnic outgroup. Adaptation to other socially marginalized groups, such as the homeless or the disabled, can be hypothesized. An expansion of the sample, already underway, and further developments will provide more evidence regarding the scale's characteristics.

NOTE

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APPENDIX

Original Version of the Allophilia Scale and Version Translated into Italian

Original version	Italian version
1. In general, I have positive attitudes about African Americans	1. In generale, ho atteggiamenti positivi nei confronti dei Rumeni
2. I respect African Americans	2. Ho rispetto per i Rumeni
3. I like African Americans	3. Mi piacciono i Rumeni
4. I feel positively toward African Americans	4. Nutro sentimenti positivi riguardo ai Rumeni
5. I am at ease around African Americans	5. Mi sento a mio agio con i Rumeni
6. I am comfortable when I hang out with African Americans	6. È per me rassicurante andare in giro con i Rumeni
7. I feel like I can be myself around African Americans	7. Sento di poter essere me stesso/a con i Rumeni
8. I feel a sense of belonging with African Americans	8. Provo un senso di appartenenza con i Rumeni
9. I feel a kinship with African Americans	9. Sento un legame di familiarità con i Rumeni
10. I would like to be more like African Americans	10. Vorrei essere più simile ai Rumeni
11. I am truly interested in understanding the points of view of African Americans	11. Sono veramente interessato a capire il punto di vista dei Rumeni
12. I am motivated to get to know African Americans better	12. Sono motivato/a a conoscere meglio i Rumeni
13. To enrich my life, I would try and make more friends who are African Americans	13. Per arricchire la mia vita, vorrei provare ad avere più amici Rumeni
14. I am interested in hearing about the experiences of African Americans	14. Sono interessato/a a sentir parlare delle esperienze dei Rumeni
15. I am impressed by African Americans	15. Sono favorevolmente colpito/a dai Rumeni
16. I feel inspired by African Americans	16. Provo fiducia nei confronti degli Rumeni
17. I am enthusiastic about African Americans	17. Sono entusiasta riguardo ai Rumeni